

A Mirage of Persistent Inequality? Comparative Educational Opportunity over the Long Haul



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May 15, 2009

The Starting Point

- Shavit and Blossfeld (1993, SB93) is a major citation hit, with Google Scholar now registering over 600 cites for the book.
- Data from 13 countries: Czech Republic, England, Germany, Hungary, Israel, Italy, Japan, the Netherlands, Poland, Sweden, Switzerland, Taiwan, the United States.
- The main conclusion is a thesis of persistent inequality of educational opportunity (IEO), measured in terms of the effects of family origins on the rates of educational transitions.

Cameron & Heckman, 1998

- Revisit the problem of dynamic selection bias in the context of Mare's sequential logit model for conditional education transitions. Propose a latent-class method to correct for dynamic selection bias.
- Criticize the arbitrary choice of effect parameters, aggravated by inattention to problems of underidentification, especially for cross-sectional data.
- When applied to US data (OCGII & NLSY), results suggest that declining IEO across transitions is not evident and depends on the choice of indices of IEO.

Design Features of SB93

- Report on OLS regressions and Mare (sequential logit) models.
- Social background indicators for IEO for most cases: father's education, father's occupation (status or EGP scheme), gender.
- Design problems of SB93:
 1. Inherent dynamic selection bias is widely acknowledged but not eliminated, so can't separate out true transition effects.
 2. Only semi-harmonized measures and models. Different chapters deal with varying # of transitions (2 to 5, seven cases with 4), hence difficult to go beyond a qualitative summary.
 3. Less obvious: effectively-small N analysis, especially when
 - breaking down into multiple cohorts,
 - examining the effects of each background variable separately &
 - at later stages of educational transition.This feature biases the main findings toward TPI—as documented by Breen, Luijkx, Muller, and Pollak (2009, AJS).

Two Motivations

- Does the thesis of persistent inequality (TPI) remain valid despite inherent dynamic selection bias?
- How is it possible that widespread educational expansion fails to reduce the influence of family background at all stages of educational transition?
 - Breen et al. have articulated an opposite thesis of nonpersistent inequality (TNI) and offered a new empirical test of TPI vs TNI for 8 European countries.
 - They found: TNI is strongly supported; the old evidence & support for TPI is misguided, largely driven by effectively small N.
- High time for a major replication of SB93's study
 - with due adjustment for bias and much larger samples,
 - a daunting task but feasible with our collaboration.

- International Stratification and Mobility File (ISMF)
 - Nationally representative samples.
 - Overlapping surveys smooth out survey effects.
 - Always: measure of father's occupation. Often: father's & mother's occupation.

- Harmonization:
 - Father's occupation: all sources recoded into ISCO68 and ISCO88, then scaled by ISEI. Range: 10-90.
 - Father's education: scaled according to level / duration. Range: 0-22 (truncated).
 - Education: organized in 7 levels, ranging from 0 No Education and 6 (Higher/Upper Tertiary).

Extract from ISMF

- Age 25-64. Cohorts born 1900-1980, coded in 10-year blocks.
- Cases with valid data on AGE, FED, FSEI and EDU.
- We have few observations in (0) No Education and (1) Incomplete Primary. Four transitions remain:
 - ED23 From Complete Primary to Lower Secondary and up.
 - ED34 From Lower Secondary to Higher Secondary and up.
 - ED45 From Higher Secondary to Lower Tertiary and up.
 - ED56 From Lower Tertiary to Upper Tertiary.

An Overview of the SB93 Samples

| Country | Authors | Sample | ISMF +=has | Cohorts | Age | Fathers' Occ | Fathers Ed | N | Transi- tions | OLS Trend |
|------------|--------------------------|---|---------------|---------------------------------|--------------|---------------------|---------------------|-------|------------------|-------------------|
| CSK | Mateju | SCSC, 1984 | + | 1918-1957, 9 years wide | | EGP5 | 4 levels | 6000 | 2 | Some |
| ENG | Kerckhoff & Trott | Oxford Study, 1972, four cohorts, men | + | 1913-1962, 10 year wide (4) | 20-59 men | SIOPS | Years | 7626 | 3 | No |
| GER | Blossfeld | GSOEP, 1984-1988 | | 1916-1965, 5 year wide (10) | | Wegener Prestige | Average duration | 4199 | 4 | No |
| HUN | Szelenyi & Aschaffenburg | SMLH, 1983 | + | 1911-1960, 10 years wide (5) | 21-72 | SIOPS | Years | 24824 | 5 | Yes, for men only |
| ISR | Shavit | Mobility Study, 1974 and Political Attitudes Study, 1988. Arab samples. | +? | 1930-1970, 10 year wide (4) | | SEI | 3 levels | 2579 | | No |
| ITA | Cobalti & Schizzerotto | ISM, 1985 | + | 1920-1961, 14 years wide (3) | | EGP, SEI | 4 levels | 4200 | 3 | Only FED |
| JAP | Treiman & Yamaguchi | Japanese Stratification Study, 1975. Men only. | + | 1906-1955, 10 years wide (5) | 20-69 men | SIOPS | Years | 2100 | 3 | Yes, FED |
| NET | De Graaf & Ganzeboom | 10 surveys, 1970-1987 | + | 1891-1960, 10 year wide (7) | >25 | SEI / EGP | 4 levels | 11244 | 3 | Yes |
| POL | Heyns & Bialecki | Social Structure and Mobility, 1987 | + | 1920-1969, 10-17 years wide (4) | 21-65, 21-60 | SIOPS, EGP9 | Years | 5434 | 4 | no |
| SWE | Jonsson | ULF, 1976-1987 | | 1902-1961, 9 year wide (7) | 26-74 | EGP6 | 5 levels | 17276 | 4 | Only Focc |
| SWI | Buchman & Charles | Career Study, 1989. Two Swiss-German cohorts born 1950 and 1960 | | 1950-1960, 3 years wide (2) | 30, 40 | SIOPS | Years | 1931 | 5 | No |
| TAI | Tsai & Chiu | Island wide survey, 1988. Men only | | 1919-1968, 11-15 years wide (3) | men | SEI | Schooling | 988 | 4 | Yes |
| USA | Hout, Raftery & Bell | GSS, Experienced Labor Force | + | 1905-1954, 10 year wide (6) | | Siegel Prestige | 5 cats | 8876 | 4 | Yes |

Sample Size Comparison: ISMF versus SB93

| | | | | | | |
|--------|--------|--------|--------|--------|--------|-------|
| CZE | ENG | GER | HUN | ISR | ITA | JAP |
| 13,068 | 10,404 | 31,518 | 83,806 | 12,714 | 36,520 | 8,473 |
| > | > | > | > | > | > | > |
| 6,000 | 7,626 | 4,199 | 24,824 | 2,579 | 4,200 | 2,100 |
| NET | POL | SWE | SWI | TAI | USA | |
| 61,756 | 76,625 | 8,532 | 5,547 | 39,977 | 57,880 | |
| > | > | < | > | > | > | |
| 11,244 | 5,434 | 17,276 | 1,931 | 988 | 8,876 | |

Analytics-1

- Like SB93, IEO here is based on logit coefficients of parental background.
- Focused on father education and SEI:
 - This focus is most directly comparable to the focus of the Blau and Duncan tradition.
 - A single measure of Total Family Effect = “Sum of partial FED & net FSEI”.
 - But also compare (a) total effect of FED & (b) partial effect of FSEI (net of FED)

Analytics-2

- All patterns are effectively “margin-free”—
 - free of systemic variation or pure noise in the marginal distributions of education and so on,
 - i.e., logit models are estimated after offsetting (as deviations from) the observed country-cohort-transition odds of making a transition.
- Explicitly test for linear trends & interactions with models of micro data.
 - As useful first-order summary of temporal trends, dramatically reducing the number of parameters.
 - Easy to visualize and conduct significance test.

Analytics-3

- Additionally, to implement the Cameron-Heckman correction for dynamic selection bias with cross-sectional data, we apply a latent-class logit regression model (LatentGold 4.0)
 - Stipulating two to three probability masses as the basis of nonparametric approximation for the stable component of unobserved heterogeneity.
 - Note that this happens to be a clever approx. to a one-dimensional continuous latent variable.
- A recent simulation study has demonstrated that the method works remarkably well in recovering true persistence of inequality using cross-sectional data (Tam 2008).

Country/Society List & Codes

- LIST

Czech Republic, England, Germany, Hungary, Israel, Italy, Japan, the Netherlands, Poland, Sweden, Switzerland, Taiwan, the United States.

- CODES

Except for USA, a case label in the figure is the first 3 letters of the name of a country/society.

Fig 1a. Consequences of Adjusting for Dynamic Selection Bias (3C) or Not (1C)

Partial FED Effect across Transitions for Oldest (0) & Youngest Cohorts (1)

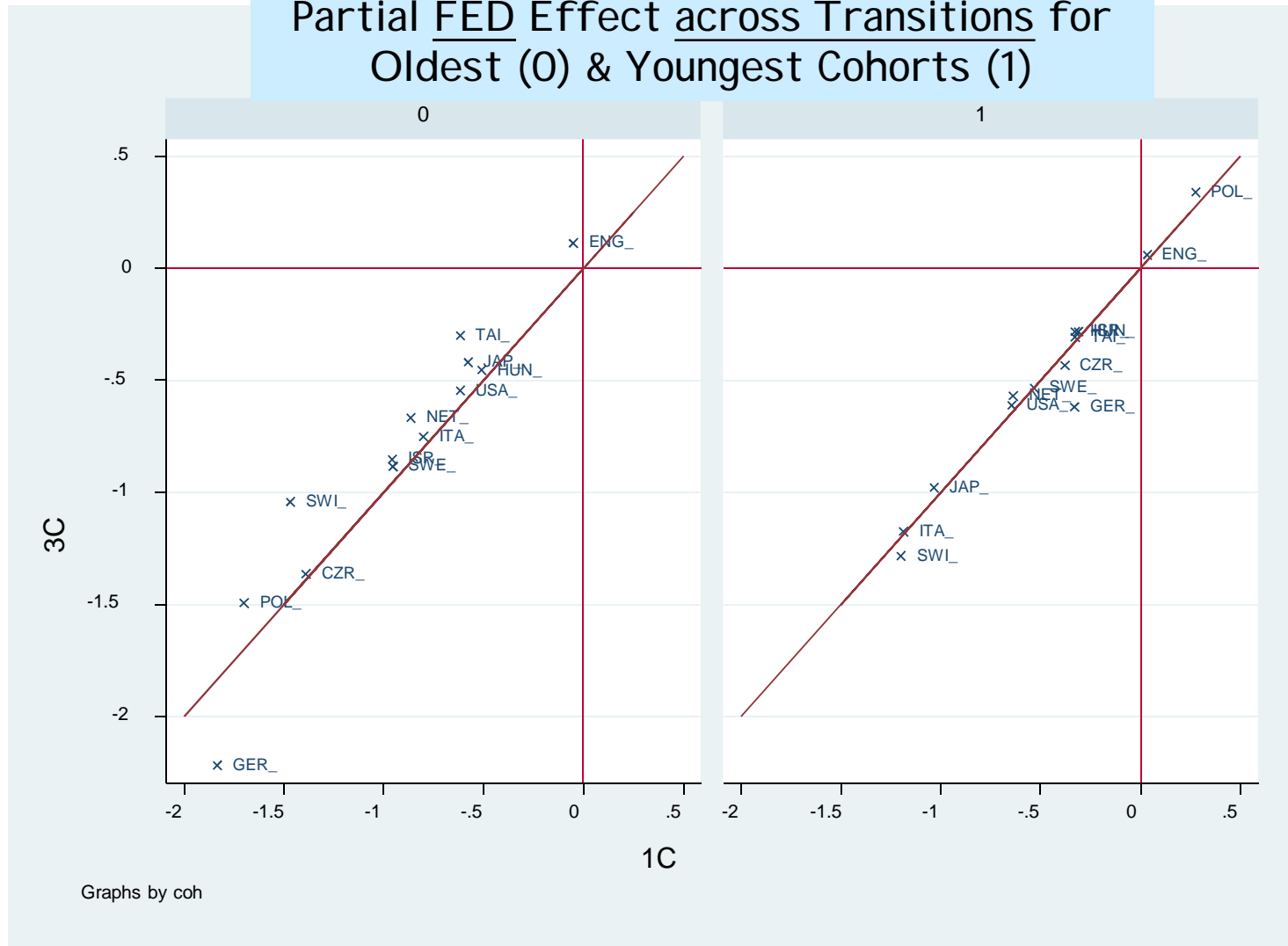


Fig 1b. Consequences of Adjusting for Dynamic Selection Bias (3C) or Not (1C)

Partial FSEI Effect across Transitions for Oldest (0) & Youngest Cohorts (1)

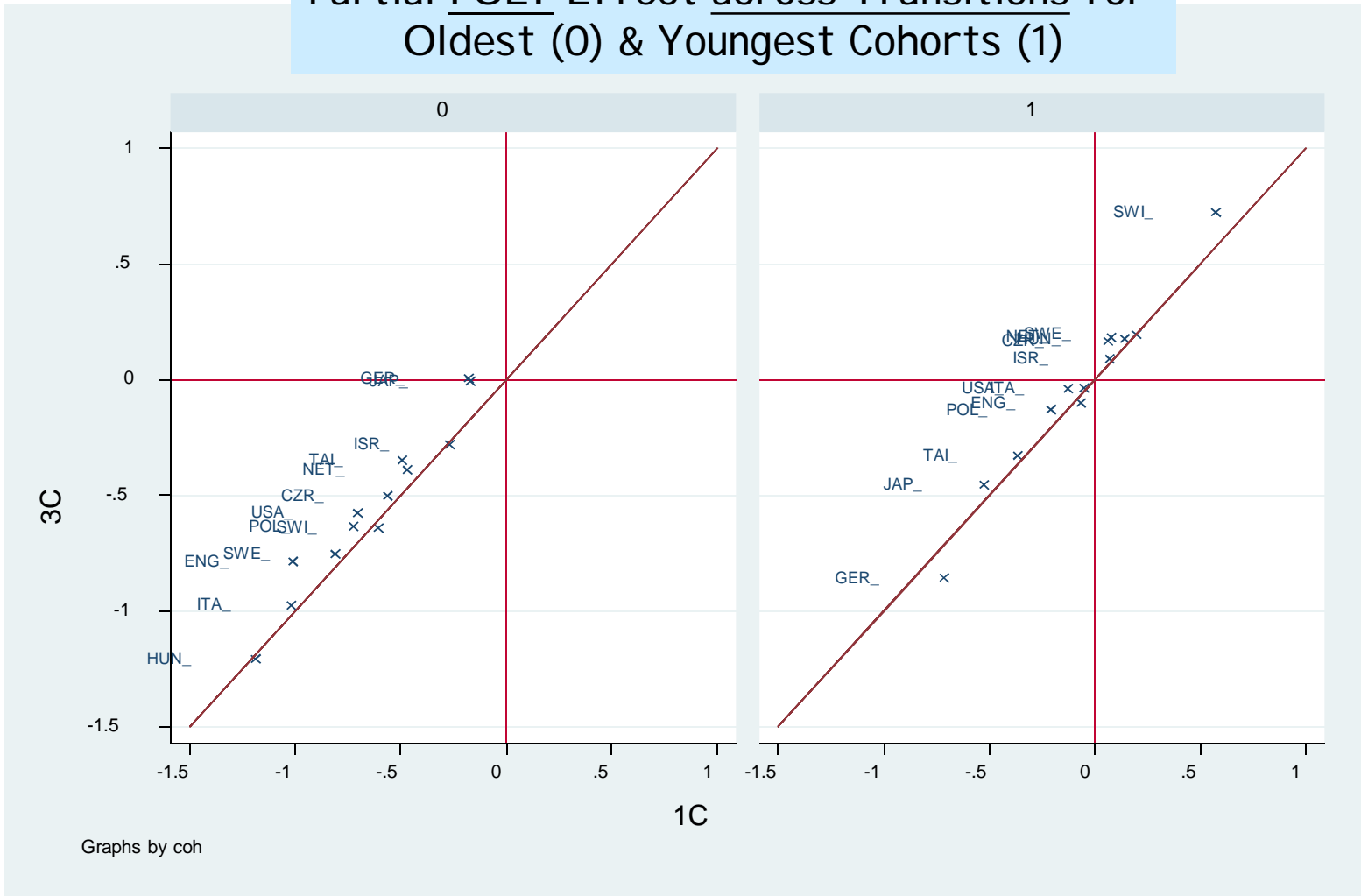
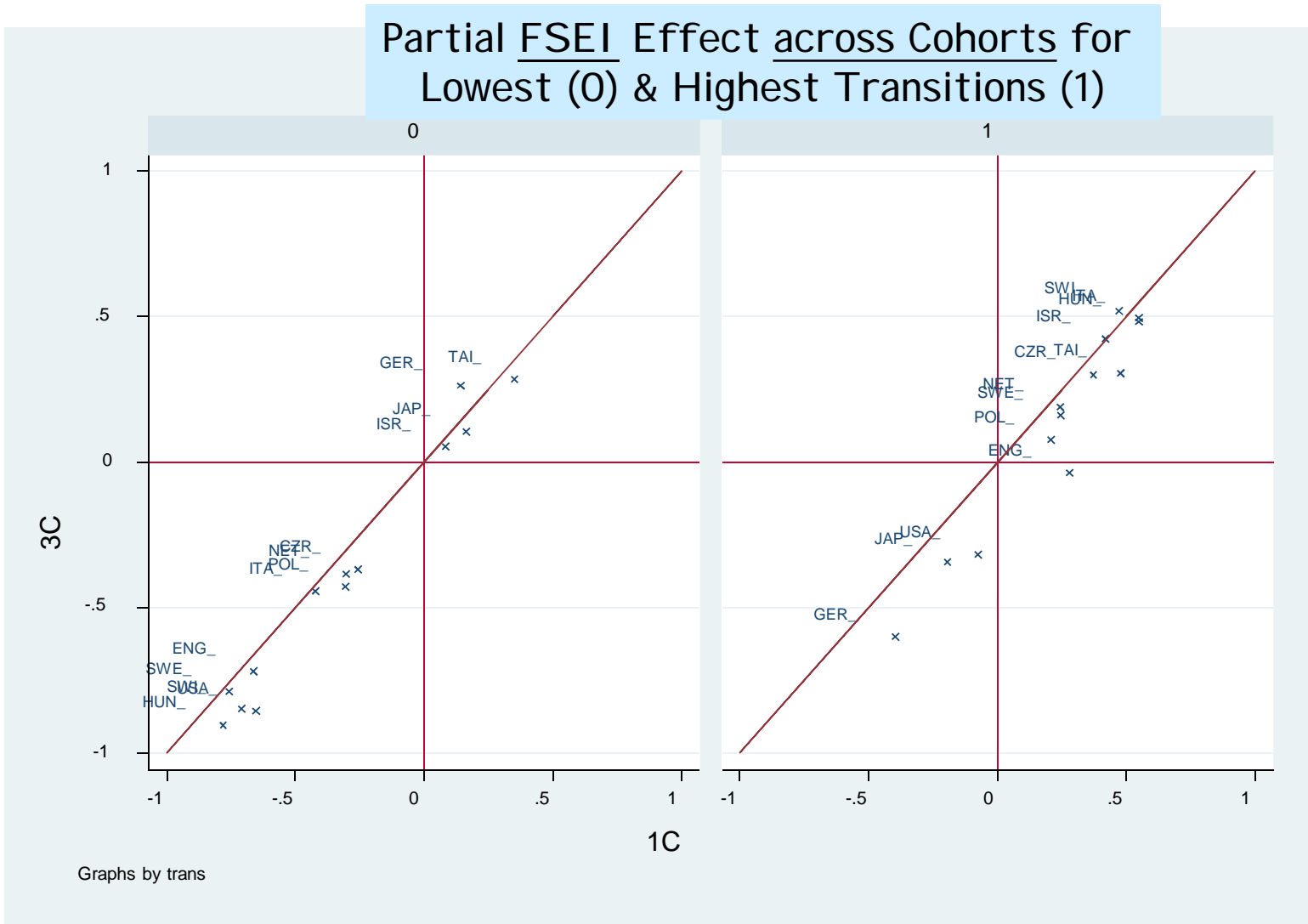


Fig 1c. Consequences of Adjusting for Dynamic Selection Bias (3C) or Not (1C)



Punch Line 1 (Adjustment for Bias)

- To our pleasant surprise, adjustment for dynamic selection bias in general does not alter any of the qualitative results; both the life-cycle and cohort trends in IEO remain intact.
 - Even though dynamic selection bias is present, the impact of the bias in the context of our 13 countries proves to be quantitatively minor and qualitatively inconsequential.

- Life-cycle dynamics (IEO across transitions): Life-cycle decline is real. The widely observed phenomenon of declining IEO from low to high educational transitions remains quantitatively strong after adjustment for dynamic selection bias.
 - That is, only a small fraction of the unadjusted decline is a statistical artifact.

The Curse of Hyper-dimensionality

- The next central finding is much harder to present: there are simply too many parameters involved.
 - Even the analysis based on the simplest specification of cohort trends & variation across 4 transitions & 13 societies results in the need to digest patterns (Fig 2a) determined by about 100 parameters.
 - Adding the nonlinear trend for the average transition (i.e. the transition experience of a representative person) brings the total number of relevant parameters to about 300.
- Our solution to the curse of dimensionality is graphical.

Fig 2a. Sum of Father Education & SEI Effects

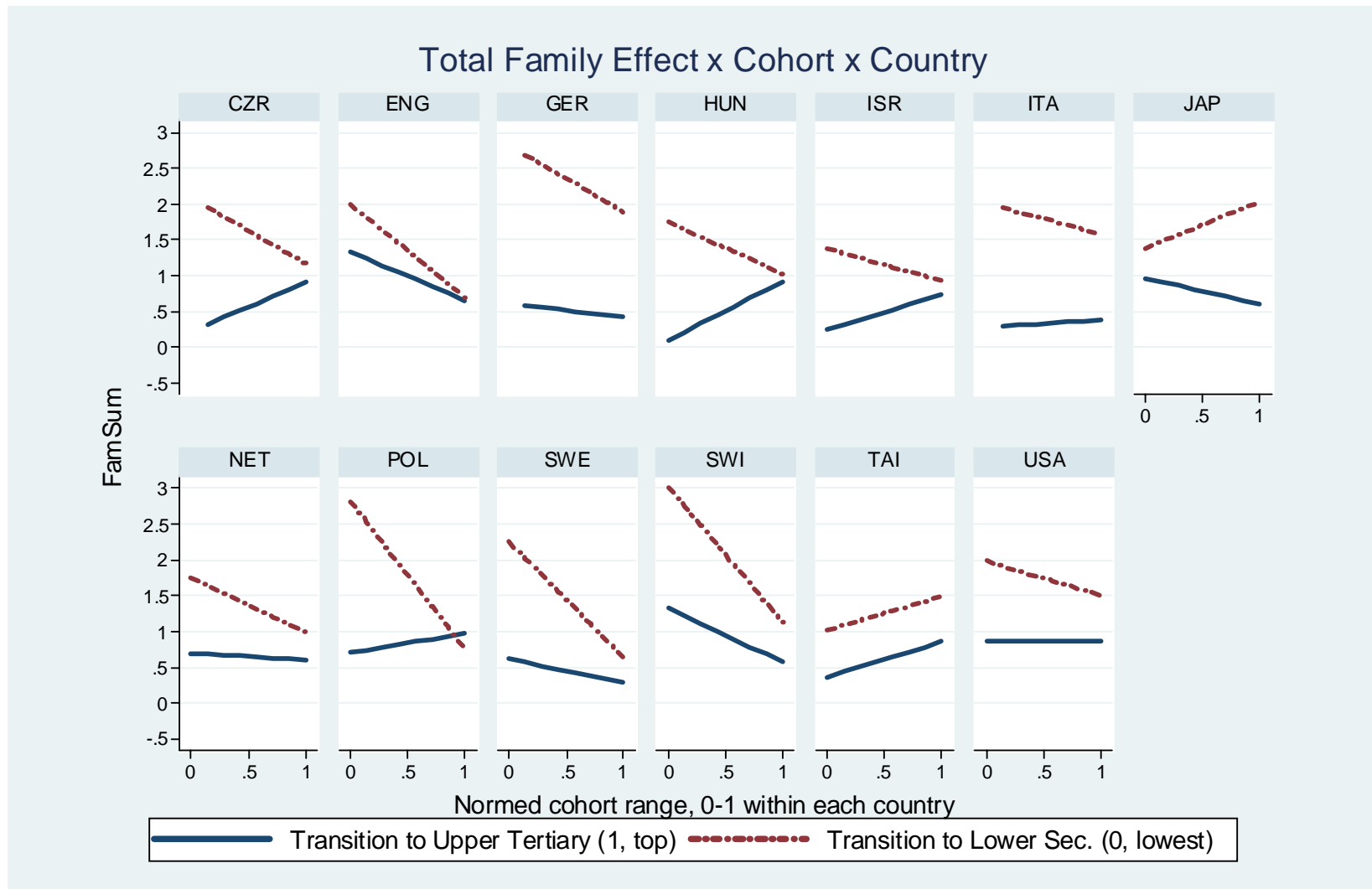


Fig 2b. The Average Transition

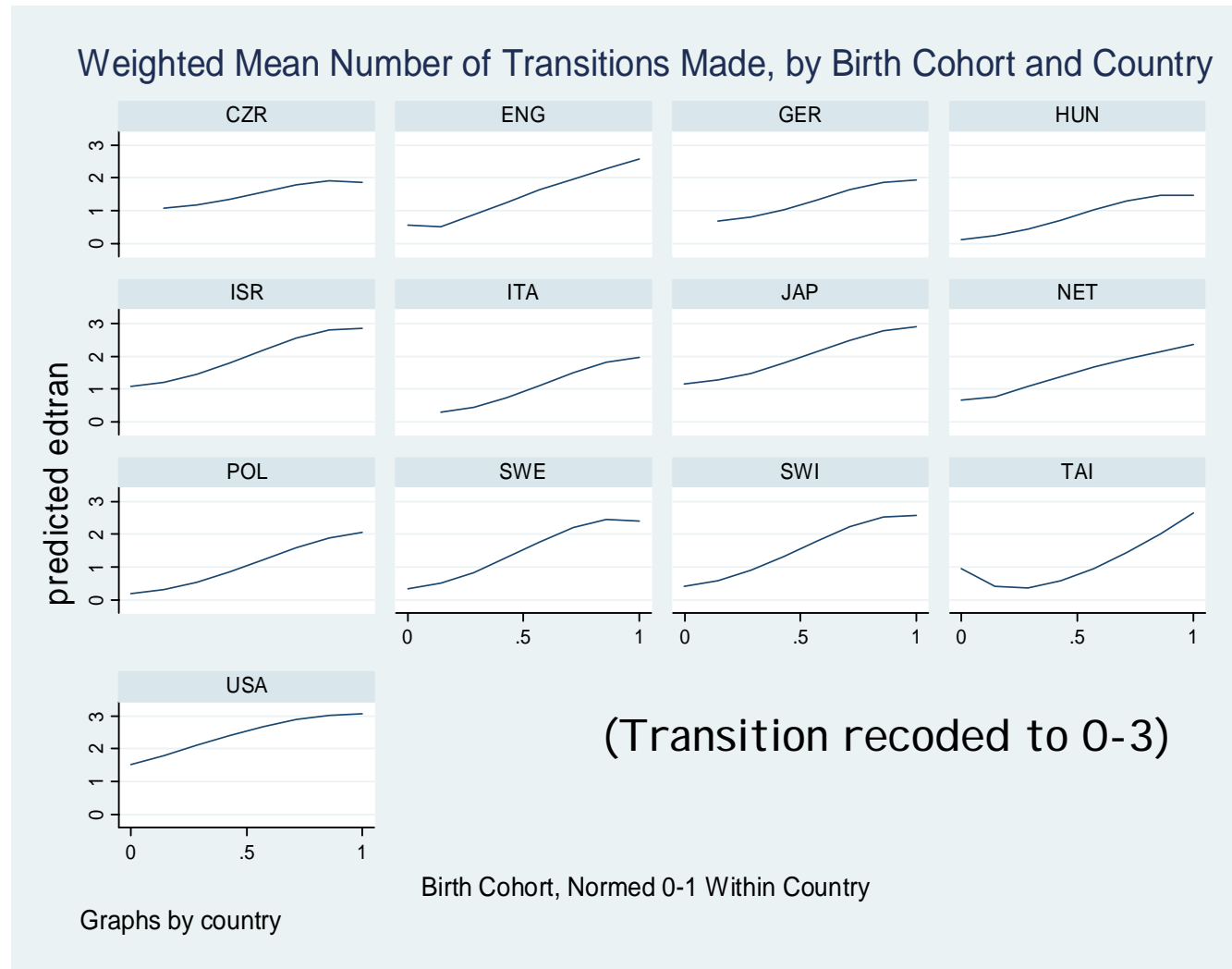
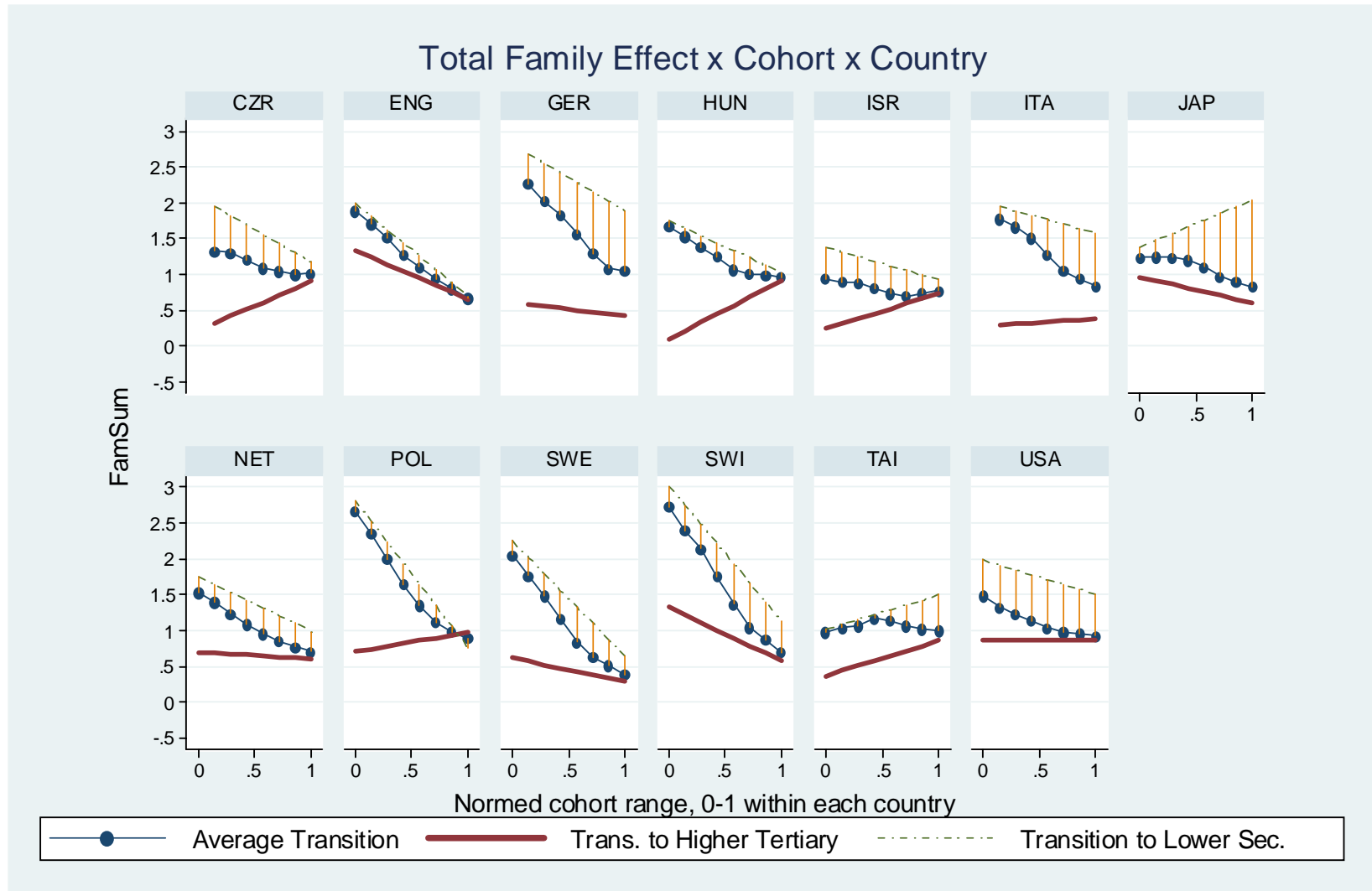


Fig 2c. Sum of Father Education & SEI Effects

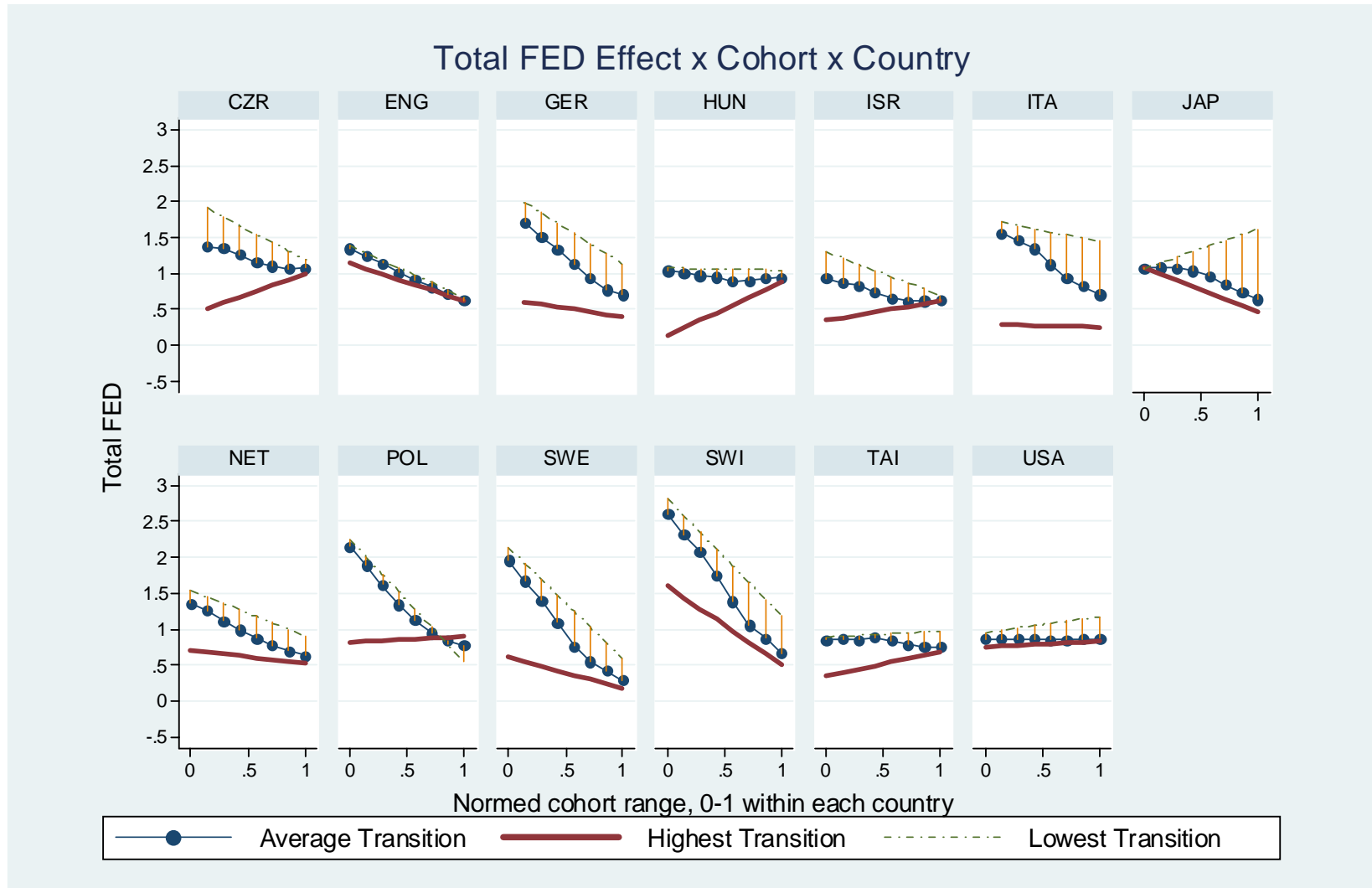


Punch Line 2 (Intercohort Trend)

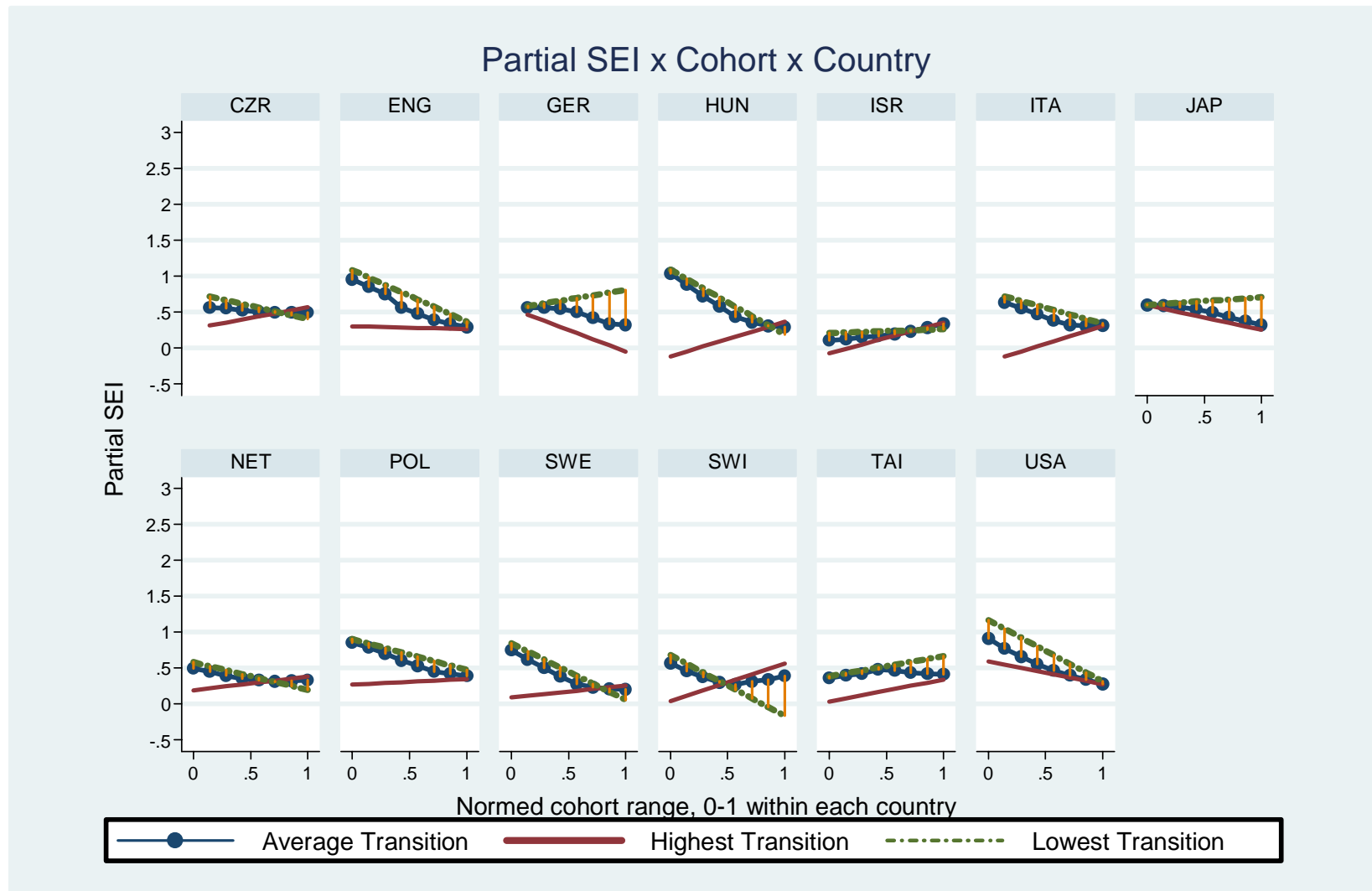
- As far as our index of total family effects is concerned, persistent inequality is hardly the norm for most societies in the twentieth century (Figure 2).
 - Specifically, pervasive long-term convergence of IEO for the highest and lowest transitions.
 - The exceptions are Japan (divergence) and Taiwan (parallel).

- For most societies, the cumulative experience of IEO has been in decline.
 - If we zoom in on the “average transition” experienced by a typical person within each cohort, the cumulative experience of IEO as a person travels from the bottom to the average transition can be represented by the shaded area.

Main Engine of IEO: Total Father Education Effect



Minor Component of IEO: Father SEI



Additional Findings

- Source of the decline: Mostly driven by declining total IEO at low transitions (note tight directional coupling of line for $\text{tran}=0$ & line for mean trans in 11 cases; and small area between the two lines for 7—CZR, ENG, HUN, NET, POL, SWE, SWI).
 - Aided by a new graphic tool, we can show that ITA, JAP, recent TAI, and GER are the only ones showing substantial role of increased attainment in lowering mean trans IEO.

- FED matters most. TOTFED, not partial FSEI, is the driver of the size and cohort trend of overall IEO.
 - When focus on the shaded area for TOTFED, can see consistent decline (not so much for HUN, TAI, USA, JAP).
 - In contrast, partial FSEI has 8 declines. Rest are quite stable.